

Does Mortgage Deregulation Increase Foreclosures? Evidence from Cleveland

Yilan Xu

Department of Economics
University of Pittsburgh

November 18, 2011

Financial Crisis and Financial Regulation

- The financial crisis of 2007 has re-ignited a debate about the impact of regulating consumer mortgage markets.
- Several scholars and prominent policymakers argue that the deregulation of mortgage lending markets has caused the large number of loan defaults and foreclosures (Warren [2007]).
- However, Ben Bernanke argues that lighter regulation of mortgage markets can spur financial innovations that broadly benefit low income households (Bernanke [2009]).

Mortgage Deregulation and Foreclosures

- Does deregulation on mortgage markets enable lenders to take advantage of uninformed borrowers?
 - If so, we would observe that deregulation would cause bad loans as a share of overall loans to increase.
- Or, does deregulation enable more credit-worthy borrowers to obtain good loans?
 - If so, we would observe that the share of bad loans following a deregulation does not increase.
- This paper uses a court-ordered repeal of home mortgage regulations in Cleveland Ohio in order to answer these questions.

Evidence from Cleveland: A Natural Experiment

- My paper examines how a court-mandated repeal of a local predatory lending law in Cleveland affected home mortgage foreclosure and origination.
- The difference-in-difference (DID) estimations indicate that following the deregulation,
 - loan foreclosures increased by 49%.
 - overall loan originations did not change.
 - loans with subprime interest rates increased by 30% and loans issued by subprime lenders increased by 40%.

Literature

- State mortgage regulations and subprime lending: Ho and Pennington-Cross [2006], Bostic et al. [2008].
- State Mortgage regulations and foreclosure rates: Ding et al. [2011]
- State foreclosure laws and home mortgage lending: Clauretie and Herzog [1990] and Pence [2006].
- Social and economics impact of foreclosures: Immergluck and Smith [2006b], Cui [2010], Immergluck and Smith [2005], Immergluck and Smith [2006a]; Schloemer et al. [2006], Mian et al. [2011].

Repeal of Cleveland Predatory Lending Law

- On November 20, 2006, the Ohio Supreme Court ruled Cleveland's predatory lending ordinances unconstitutional.
- The court rule sided with American Financial Services Association (AFSA), a national organization that challenged the city law shortly after the enactment in 2002.
- Cleveland law regulated the loans secured by owner-occupied residential properties located within the city limit of Cleveland.
- This deregulation makes Cleveland a desirable subject of a natural experiment to study the impact of lending deregulation.

Federal, State and Local Predatory Lending Laws

- Following Home Ownership and Equity Protection Act (HOEPA) enacted by Congress, most state and local predatory lending laws are also defined in two parts:
 - The first part defines the coverage of the laws by product types, interest rate triggers and fees triggers.
 - The second part imposes certain disclosures and lending restrictions on the covered loans.
- Typically, the state and local laws enhance both the coverage and the restrictions of the federal law.

A Comparison of Ohio Law and Cleveland Law

- Covered Loan Types:
 - Ohio: home equity loans;
 - Cleveland: all home loans, including home-purchase loans.
- Interest rate triggers (first-lien):
 - Ohio: 8 percentage points above the treasury rate;
 - Cleveland: 4.5 and 8 percentage points above the treasury rate.

A Comparison of Ohio Law and Cleveland Law

In addition to the restrictions implemented by state law, the Cleveland Ordinance imposed restrictions on:

- loan flipping, balloon payments, negative amortization,
- an increased interest rate on default, advance payments, mandatory arbitration, prepayment penalties, financing of credit insurance,
- lending without counseling, lending without due regard to prepayment, and payments to home improvement contractors under certain circumstances.

Difference-in-difference Identification

- A natural experiment
 - Treatment group: census tracts in Cleveland.
 - Control group: census tracts in the suburban municipalities.
 - Treatment: deregulation.
- The causal inference of the DID estimation assumes common time trends in the absence of the deregulation.
- The DID method identifies the impact of deregulation as the deviation between time trends of Cleveland and the suburban municipalities following deregulation.

Data

- The Loan Origination and Foreclosure Matched Data of Cuyahoga County: loans made during 2005-2008, foreclosures by the end of December, 2009.
- The sample included home purchase loans for 1- to 4-family housing units secured by owner-occupied housing properties in Cuyahoga county.
- Early foreclosure is defined by foreclosure complaints filed by lenders within 30 months after origination.
- The constructed longitudinal data set includes loan counts at tract-month level during June 2006 – May 2007.
 - 458 tracts, 5496 monthly observations;
 - 6 months before and after the deregulation.

Panel Poisson Estimation

$$E(y_{it}|X) = \exp(\theta_i + \eta_t + \alpha_1\delta^c + \alpha_2\gamma^R + \alpha_3(\delta^c \times \gamma^R)) \quad (1)$$

- y_{it} is the loan count for census tract i in month t .
- θ_i is the tract fixed effects, and η_t is the month fixed effect.
- δ^c is a dummy for tracts in Cleveland.
- γ^R is a dummy for the periods after the repeal.
- α_3 measures the treatment effect.

Panel Poisson Estimation

$$E(y_{it}|X) = \exp(\theta_i + \eta_t + \alpha_1 \delta^c + \alpha_2 \gamma^R + \alpha_3 (\delta^c \times \gamma^R)) \quad (2)$$

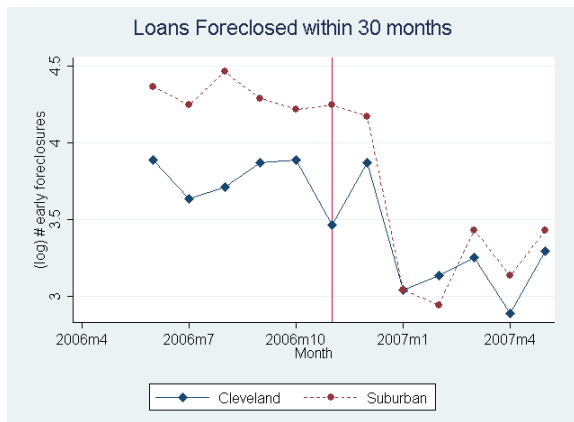
- α_3 can be interpreted as the log odds ratio.

$$\alpha_3 = \log \frac{E(y_{Cleveland,after}|X)}{E(y_{Cleveland,before}|X)} / \frac{E(y_{Subs,after}|X)}{E(y_{Subs,before}|X)} \quad (3)$$

- Panel Poisson specification assumes the treatment group and the control group have proportional changes in outcome variables over time in the absence of the treatment.
- $\alpha_3 > 0 \implies \text{odds ratio} > 1 \implies$ the event is more likely in Cleveland after deregulation.

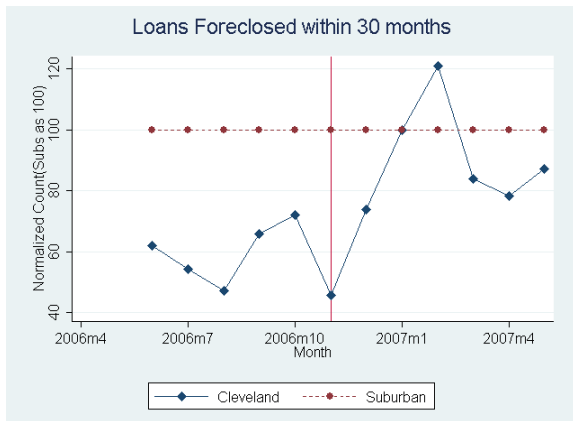
Difference-in-difference: Foreclosures

Figure 1: Time Trends, June 2006 – May 2007



Difference-in-difference: Foreclosures

Figure 2: Time Trends, June 2006 – May 2007



Difference-in-difference: Foreclosures

Table 1: Impact of Deregulation on Foreclosures

| VARIABLES | (1) Baseline |
|------------------|-------------------|
| Cleveland×Repeal | 0.40*** (0.14) |
| Observations | 4,200 |
| Number of tracts | 350 |
| Tract FE | YES |
| Month FE | YES |
| Control | NO |

Notes: The point estimate of 0.4 translates into an odds ratio of 1.49 and implies a 49% increase in the number of loans with early foreclosures in Cleveland after the repeal.

Difference-in-difference: Foreclosures

Table 1: Impact of Deregulation on Foreclosures

| VARIABLES | (1) Baseline | (2) Non-preempted |
|------------------|-------------------|----------------------|
| Cleveland×Repeal | 0.40*** (0.14) | 0.47*** (0.16) |
| Observations | 4,200 | 2,681 |
| Number of tracts | 350 | 313 |
| Tract FE | YES | YES |
| Month FE | YES | YES |
| Control | NO | NO |

Notes: The point estimate of 0.4 translates into an odds ratio of 1.49 and implies a 49% increase in the number of loans with early foreclosures in Cleveland after the repeal.

Difference-in-difference: Foreclosures

Table 1: Impact of Deregulation on Foreclosures

| VARIABLES | (1) Baseline | (2) Non-preempted | (3) b=3 months |
|------------------|-------------------|----------------------|-------------------|
| Cleveland×Repeal | 0.40*** (0.14) | 0.47*** (0.16) | 0.36** (0.18) |
| Observations | 4,200 | 2,681 | 1,560 |
| Number of tracts | 350 | 313 | 260 |
| Tract FE | YES | YES | YES |
| Month FE | YES | YES | YES |
| Control | NO | NO | NO |

Notes: The point estimate of 0.4 translates into an odds ratio of 1.49 and implies a 49% increase in the number of loans with early foreclosures in Cleveland after the repeal.

Difference-in-difference: Foreclosures

Table 1: Impact of Deregulation on Foreclosures

| VARIABLES | (1) Baseline | (2) Non-preempted | (3) b=3 months | (4) Border |
|------------------|-------------------|----------------------|-------------------|-----------------|
| Cleveland×Repeal | 0.40*** (0.14) | 0.47*** (0.16) | 0.36** (0.18) | 0.44* (0.24) |
| Observations | 4,200 | 2,681 | 1,560 | 1,332 |
| Number of tracts | 350 | 313 | 260 | 111 |
| Tract FE | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES |
| Control | NO | NO | NO | NO |

Notes: The point estimate of 0.4 translates into an odds ratio of 1.49 and implies a 49% increase in the number of loans with early foreclosures in Cleveland after the repeal.

Difference-in-difference: Foreclosures

Table 1: Impact of Deregulation on Foreclosures

| VARIABLES | (1) Baseline | (2) Non-preempted | (3) b=3 months | (4) Border | (5) 24-month |
|------------------|-------------------|----------------------|-------------------|-----------------|-------------------|
| Cleveland×Repeal | 0.40*** (0.14) | 0.47*** (0.16) | 0.36** (0.18) | 0.44* (0.24) | 0.49*** (0.16) |
| Observations | 4,200 | 2,681 | 1,560 | 1,332 | 3,804 |
| Number of tracts | 350 | 313 | 260 | 111 | 317 |
| Tract FE | YES | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES | YES |
| Control | NO | NO | NO | NO | NO |

Notes: The point estimate of 0.4 translates into an odds ratio of 1.49 and implies a 49% increase in the number of loans with early foreclosures in Cleveland after the repeal.

Falsifications: Timing and Locality

Table 2: Falsifications for Early Foreclosures

| | (1) Nov. 2005 6 month |
|------------------|-----------------------------|
| City×Repeal | -0.13 (0.11) |
| Observations | 4,524 |
| Number of tracts | 377 |
| Tract FE | YES |
| Month FE | YES |
| Control | NO |

Falsifications: Timing and Locality

Table 2: Falsifications for Early Foreclosures

| | (1) Nov. 2005 6 month | (2) May 2006 6 month |
|------------------|-----------------------------|----------------------------|
| City×Repeal | -0.13 (0.11) | -0.17 (0.12) |
| Observations | 4,524 | 4,320 |
| Number of tracts | 377 | 360 |
| Tract FE | YES | YES |
| Month FE | YES | YES |
| Control | NO | NO |

Falsifications: Timing and Locality

Table 2: Falsifications for Early Foreclosures

| | (1) Nov. 2005 6 month | (2) May 2006 6 month | (3) Inner Subs 6 month |
|------------------|-----------------------------|----------------------------|------------------------------|
| City×Repeal | -0.13 (0.11) | -0.17 (0.12) | 0.02 (0.20) |
| Observations | 4,524 | 4,320 | 2,544 |
| Number of tracts | 377 | 360 | 212 |
| Tract FE | YES | YES | YES |
| Month FE | YES | YES | YES |
| Control | NO | NO | NO |

Falsifications: Timing and Locality

Table 2: Falsifications for Early Foreclosures

| | (1) Nov. 2005 6 month | (2) May 2006 6 month | (3) Inner Subs 6 month | (4) Pittsburgh 6 month |
|------------------|-----------------------------|----------------------------|------------------------------|------------------------------|
| City×Repeal | -0.13 (0.11) | -0.17 (0.12) | 0.02 (0.20) | 0.32 (0.24) |
| Observations | 4,524 | 4,320 | 2,544 | 2,412 |
| Number of tracts | 377 | 360 | 212 | 201 |
| Tract FE | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES |
| Control | NO | NO | NO | NO |

Falsifications: Timing and Locality

Table 2: Falsifications for Early Foreclosures

| | (1) Nov. 2005 6 month | (2) May 2006 6 month | (3) Inner Subs 6 month | (4) Pittsburgh 6 month | (5) Pittsburgh 3 month |
|------------------|-----------------------------|----------------------------|------------------------------|------------------------------|------------------------------|
| City×Repeal | -0.13 (0.11) | -0.17 (0.12) | 0.02 (0.20) | 0.32 (0.24) | -0.02 (0.39) |
| Observations | 4,524 | 4,320 | 2,544 | 2,412 | 756 |
| Number of tracts | 377 | 360 | 212 | 201 | 126 |
| Tract FE | YES | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES | YES |
| Control | NO | NO | NO | NO | NO |

Falsifications: Housing Bubble and Bad Economy

Table 3: Social and Economic Conditions Before and After the Repeal

| | (1) Sales Price |
|---------------------------|--------------------------|
| Cleveland \times Repeal | -7,537.35 (11,087.85) |
| Observations | 28,006 |
| Tract FE | YES |
| Month FE | YES |
| Control | NO |

Falsifications: Housing Bubble and Bad Economy

Table 3: Social and Economic Conditions Before and After the Repeal

| | (1) Sales Price | (2) Sales Price |
|--------------------|--------------------------|--------------------------|
| Cleveland × Repeal | -7,537.35 (11,087.85) | -6,678.26 (11,032.27) |
| Observations | 28,006 | 28,006 |
| Tract FE | YES | YES |
| Month FE | YES | YES |
| Control | NO | YES |

Falsifications: Housing Bubble and Bad Economy

Table 3: Social and Economic Conditions Before and After the Repeal

| | (1) | (2) | (3) |
|---------------------------|--------------------------|--------------------------|------------------|
| | Sales Price | Sales Price | Foreclosed Homes |
| Cleveland \times Repeal | -7,537.35 (11,087.85) | -6,678.26 (11,032.27) | -0.07 (0.13) |
| Observations | 28,006 | 28,006 | 4,032 |
| Tract FE | YES | YES | YES |
| Month FE | YES | YES | YES |
| Control | NO | YES | NO |

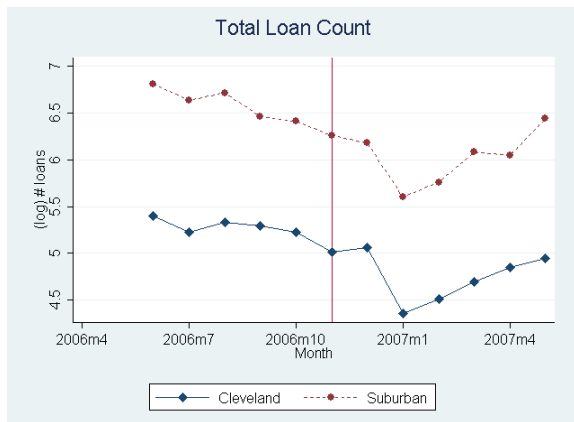
Foreclosure: By Loan Types and Lender Types

Table 4: Early Foreclosures by Types

| | (1) All | (2) Subprime | (3) Prime | (4) Subprime | (5) Prime |
|--------------------|-------------------|-------------------|----------------|-----------------|-----------------|
| Cleveland × Repeal | 0.40*** (0.14) | 0.50*** (0.17) | 0.39 (0.27) | 0.44* (0.23) | 0.32* (0.18) |
| Observations | 4,200 | 3,528 | 1,722 | 2,664 | 3,444 |
| Number of tracts | 350 | 294 | 186 | 222 | 287 |
| Tract FE | YES | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES | YES |
| Control | NO | NO | NO | NO | NO |

Difference-in-difference: Originations

Figure 3: Total Loan Count, June 2006 – May 2007



Difference-in-difference: Originations

Table 5: Impact of Deregulation on Loan Count

| VARIABLES | (1) Baseline |
|---------------------------|-----------------|
| Cleveland \times Repeal | 0.02 (0.05) |
| Observations | 5,496 |
| Number of tracts | 458 |
| Tract FE | YES |
| Month FE | YES |
| Control | NO |

Difference-in-difference: Originations

Table 5: Impact of Deregulation on Loan Count

| VARIABLES | (1) Baseline | (2) Non-preempted |
|---------------------------|-----------------|----------------------|
| Cleveland \times Repeal | 0.02 (0.05) | 0.03 (0.06) |
| Observations | 5,496 | 3,517 |
| Number of tracts | 458 | 436 |
| Tract FE | YES | YES |
| Month FE | YES | YES |
| Control | NO | NO |

Difference-in-difference: Originations

Table 5: Impact of Deregulation on Loan Count

| VARIABLES | (1) Baseline | (2) Non-preempted | (3) b=3 months |
|--------------------|-----------------|----------------------|-------------------|
| Cleveland × Repeal | 0.02 (0.05) | 0.03 (0.06) | 0.01 (0.08) |
| Observations | 5,496 | 3,517 | 2,622 |
| Number of tracts | 458 | 436 | 437 |
| Tract FE | YES | YES | YES |
| Month FE | YES | YES | YES |
| Control | NO | NO | NO |

Difference-in-difference: Originations

Table 5: Impact of Deregulation on Loan Count

| VARIABLES | (1) Baseline | (2) Non-preempted | (3) b=3 months | (4) Border |
|---------------------------|-----------------|----------------------|-------------------|-----------------|
| Cleveland \times Repeal | 0.02 (0.05) | 0.03 (0.06) | 0.01 (0.08) | -0.04 (0.09) |
| Observations | 5,496 | 3,517 | 2,622 | 1,560 |
| Number of tracts | 458 | 436 | 437 | 130 |
| Tract FE | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES |
| Control | NO | NO | NO | NO |

Loan Origination: By Loan Types and Lender Types

Table 6: Loan Count by Types

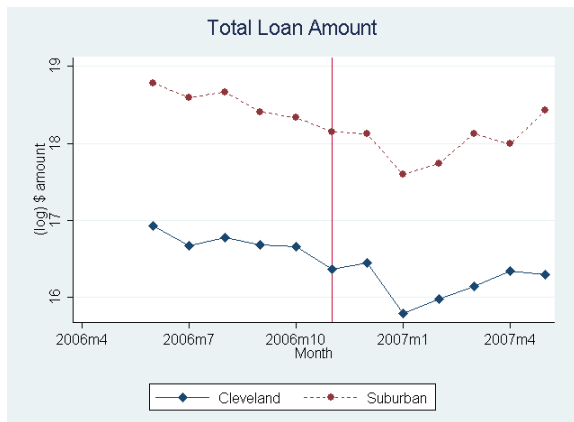
| | (1) All | (2) Subprime | (3) Prime | (4) Subprime | (5) Prime |
|--------------------|----------------|-------------------|----------------|------------------|----------------|
| Cleveland × Repeal | 0.02 (0.05) | 0.27*** (0.09) | 0.02 (0.07) | 0.34** (0.16) | 0.06 (0.06) |
| Observations | 5,496 | 4,980 | 5,148 | 4,128 | 5,412 |
| Number of tracts | 458 | 415 | 429 | 344 | 451 |
| Tract FE | YES | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES | YES |
| Control | NO | NO | NO | NO | NO |

A coefficient of 0.27 translates into an odds ratio of 1.31 and implies a 30% increase in the subprime loans.

A coefficient of 0.34 translates into an odds ratio of 1.40 and implies a 40% increase in the loans made by subprime lenders.

Difference-in-difference: Loan Amount

Figure 4: Total Loan Amount, June 2006 – May 2007



Difference-in-difference: Loan Amount

Table 7: Impact of Deregulation on Loan Amount

| VARIABLES | (1) Baseline | (2) Non-preempted | (3) b=3 months | (4) Border |
|--------------------|-----------------|----------------------|-------------------|-----------------|
| Cleveland × Repeal | -0.03 (0.06) | -0.01 (0.06) | -0.01 (0.09) | -0.07 (0.13) |
| Observations | 5,496 | 3,517 | 2,622 | 1,560 |
| Number of tracts | 458 | 436 | 437 | 130 |
| Tract FE | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES |
| Control | NO | NO | NO | NO |

Loan Amount: By Loan Types and Lender Types

Table 8: Loan Amount by Types

| | (1) All | (2) Subprime | (3) Prime | (4) Subprime | (5) Prime |
|--------------------|-----------------|------------------|-----------------|-----------------|----------------|
| Cleveland × Repeal | -0.03 (0.06) | 0.23** (0.10) | -0.01 (0.08) | 0.34* (0.18) | 0.00 (0.07) |
| Observations | 5,496 | 4,980 | 5,148 | 4,128 | 5,412 |
| Number of tracts | 458 | 415 | 429 | 344 | 451 |
| Tract FE | YES | YES | YES | YES | YES |
| Month FE | YES | YES | YES | YES | YES |
| Control | NO | NO | NO | NO | NO |

Conclusions

- This paper uses a court mandated repeal of predatory lending law in Cleveland, OH as a natural experiment to study the impact of deregulation on credit flow and loan quality.
- Empirical results indicate that deregulation caused a 49 percent increase in early foreclosures while it did not increase total loan volume, implying that bad loans as a share of overall loans increased substantially.
- Deregulation also increased the number of high-interest loans by 30 percent, increased loans made by subprime lenders by 40 percent.

Conclusions

- The robustness checks and falsification tests provide evidence that the increased early foreclosures are not driven by the definition of early foreclosure or the seasonal pattern of foreclosure.
- Moreover, the social and economic environment in which the loans are made, and the shock from the subprime crisis are not the reasons for the increased foreclosures after deregulation.
- These results suggest that the Cleveland predatory lending law, without reducing the credit supply, implemented restrictions under which the originated loans would survive longer.